

ECSE 305A: Probability and Random Signals I  
 Problem Set 4  
 solutions

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1. (a)  $X$  is a random variable of mixed type since it is continuous except for discontinuity at 0 and at 1.

(b)

$$P(X \geq 5) = 1 - P(X < 5) = 1 - F(5^-) = 0$$

$$P(X < 0) = F(0^-) = 0$$

$$P(X \leq 0) = F(0) = \frac{1}{4}$$

$$P\left(\frac{1}{4} \leq X < 1\right) = F(1^-) - F\left(\frac{1}{4}^-\right) = \frac{3}{16}$$

$$P\left(\frac{1}{4} \leq X \leq 1\right) = F(1) - F\left(\frac{1}{4}^-\right) = \frac{11}{16}$$

$$P\left(X > \frac{1}{2}\right) = 1 - P\left(X \leq \frac{1}{2}\right) = 1 - F\left(\frac{1}{2}\right) = \frac{5}{8}$$

2. (a) It can be seen that CDF  $F(x)$  is flat except for a finite number of jumps, so  $X$  is a discrete RV.

(b)

$$p(x) = \begin{cases} x < -2 & 0 \\ x = -2 & F(-2) - F(-2^-) = 1/4 - 0 = 1/4 \\ x = -1 & F(-1) - F(-1^-) = 1/4 - 1/4 = 0 \\ x = 0 & F(0) - F(0^-) = 1/2 - 1/4 = 1/4 \\ x = 1 & F(1) - F(1^-) = 3/4 - 1/2 = 1/4 \\ x = 2 & F(2) - F(2^-) = 1 - 3/4 = 1/4 \\ x > 2 & 0 \end{cases}$$

(c) Assuming  $Y = X^2$  then  $R_y = \{0, 1, 4\}$

$$p(y) = \begin{cases} y < 0 & 0 \\ y = 0 & \sum_{x^2=0} p(x) = p(x=0) = 1/4 \\ y = 1 & \sum_{x^2=1} p(x) = p(x=1) + p(x=-1) = 1/4 + 0 = 1/4 \\ y = 4 & \sum_{x^2=4} p(x) = p(x=2) + p(x=-2) = 1/4 + 1/4 = 1/2 \\ y > 4 & 0 \end{cases}$$

3. Let  $p$  be the probability function of  $X$  and  $F$  be its distribution function. We have

$$p(i) = \left(\frac{5}{6}\right)^{i-1} \left(\frac{1}{6}\right), \quad i = 1, 2, 3, \dots$$

$F(x) = 0$  for  $x < 1$ . If  $x \geq 1$ , for some positive integer  $n$ ,  $n \leq x < n+1$ , and we have that

$$\begin{aligned} F(x) &= \sum_{i=1}^n \left(\frac{5}{6}\right)^{i-1} \left(\frac{1}{6}\right) \\ &= \frac{1}{6} * \frac{1 - \left(\frac{5}{6}\right)^n}{1 - \frac{5}{6}} \\ &= 1 - \left(\frac{5}{6}\right)^n \end{aligned}$$

Hence

$$y = \begin{cases} 0 & \text{if } x < 1 \\ 1 - \left(\frac{5}{6}\right)^n & \text{if } n \leq x < n+1, \quad n = 1, 2, 3, \dots \end{cases}$$

4.

$$\begin{aligned} \sum_{k=0}^{\infty} P(X > k) &= P(X > 0) & + P(X > 1) & + P(X > 2) + \dots \\ &= P(X = 1) & + P(X = 2) & + P(X = 3) + \dots \\ & & + P(X = 2) & + P(X = 3) + \dots \\ & & & + P(X = 3) + \dots \\ &= 1 \cdot P(X = 1) & + 2 \cdot P(X = 2) & + 3 \cdot P(X = 3) + \dots \\ &= E[X]. \end{aligned}$$

5. There are two ways to solve this problem:

(a)

$$\begin{aligned} E(Y) &= \sum_{k=0}^{\infty} Y * p_k \\ &= \sum_{k=1}^{\infty} (k-1)p_k \\ &= \sum_{k=0}^{\infty} (k-1)p_k + p_0 \\ &= E(X) - 1 + p_0 \end{aligned}$$

$$\begin{aligned}
E(Y^2) &= \sum_{k=0}^{\infty} Y^2 * p_k \\
&= \sum_{k=1}^{\infty} (k-1)^2 p_k \\
&= \sum_{k=0}^{\infty} (k-1)^2 p_k - p_0 \\
&= \sum_{k=0}^{\infty} (k^2 - 2k + 1) p_k - p_0 \\
&= E(X^2) - 2E(X) + 1 - p_0
\end{aligned}$$

(b) (Optional) Define Moment Function as:

$$\Gamma(z) = E(z^n) = \sum_{n=-\infty}^{\infty} p_n z^n$$

Differentiating  $\Gamma(z)$   $k$  times, we obtain

$$\Gamma^{(k)}(z) = E[n(n-1)(n-2)\cdots(n-k+1)z^{n-k}]$$

With  $z = 1$ , this yields

$$\Gamma^{(k)}(1) = E[n(n-1)(n-2)\cdots(n-k+1)]$$

In particular, that  $\Gamma(1) = 1$ ,  $\Gamma'(1) = E(n)$  and  $\Gamma''(1) = E(n^2) - E(n)$

In this problem, it is obvious that  $\Gamma_y(z) = p_0 + z^{-1}[\Gamma_x(z) - p_0]$ .

Therefore

$$\Gamma'_y(z) = z^{-1}(z\Gamma'_x(z) - \Gamma_x(z) + p_0)$$

and

$$\Gamma''_y(z) = z^{-1}(z^2\Gamma''_x(z) - z\Gamma'_x(z) + \Gamma_x(z) - p_0)$$

Let  $z = 1$  and we have  $\Gamma(1) = 1$ ,  $\Gamma'(1) = E(n)$  and  $\Gamma''(1) = E(n^2) - E(n)$ . Therefore

$$E(Y) = E(X) - 1 + p_0, \quad E(Y^2) = E(X^2) - 2E(X) + 1 - p_0$$

6. Let  $X$  be the number of children they should continue to have until they have one of each sex. Let  $B$  be the event that the first child is a boy. Then

$$\begin{aligned}
P(X = i) &= P(X = i|B)P(B) + P(X = i|B^c)P(B^c) \\
&= \left(\frac{1}{2}\right)^{i-2}\left(\frac{1}{2}\right) \cdot \frac{1}{2} + \left(\frac{1}{2}\right)^{i-2}\left(\frac{1}{2}\right) \cdot \frac{1}{2} \\
&= \left(\frac{1}{2}\right)^{i-1}
\end{aligned}$$

So

$$E(X) = \sum_{i=2}^{\infty} i\left(\frac{1}{2}\right)^{i-1} = -1 + \sum_{i=1}^{\infty} i\left(\frac{1}{2}\right)^{i-1} = 3$$

7. (a) Let  $X$  be the number of trials required to open the door. Clearly,

$$P(X = x) = \left(1 - \frac{1}{n}\right)^{x-1} \frac{1}{n}, \quad x = 1, 2, 3, \dots$$

Thus

$$\begin{aligned} E(X) &= \sum_{x=1}^{\infty} x \left(1 - \frac{1}{n}\right)^{x-1} \frac{1}{n} \\ &= \frac{1}{n} \frac{1}{\left[1 - \left(1 - \frac{1}{n}\right)\right]^2} \\ &= n \end{aligned}$$

To calculate  $Var(X)$ , first we find  $E(x^2)$ . We have

$$\begin{aligned} E(X^2) &= \sum_{x=1}^{\infty} x^2 \left(1 - \frac{1}{n}\right)^{x-1} \frac{1}{n} \\ &= \frac{1}{n} \sum_{x=1}^{\infty} x^2 \left(1 - \frac{1}{n}\right)^{x-1} \\ &= \frac{1}{n} \frac{1 + \left(1 - \frac{1}{n}\right)}{\left[1 - \left(1 - \frac{1}{n}\right)\right]^3} \\ &= 2n^2 - n \end{aligned}$$

Therefore

$$Var(X) = (2n^2 - n) - n^2 = n(n - 1)$$

(b) Let  $A_i$  be the event that on the  $i$ th trial the door opens. Let  $X$  be the number of trials required to open the door. Then

$$\begin{aligned} P(X = k) &= P[A_1^c \cap A_2^c \cap \dots \cap A_{k-1}^c \cap A_k] \\ &= P(A_k | A_1^c \cap A_2^c \cap \dots \cap A_{k-1}^c) * \\ &\quad P(A_{k-1}^c | A_1^c \cap A_2^c \cap \dots \cap A_{k-2}^c) * \\ &\quad \dots * P(A_2^c | A_1^c) * P(A_1) \\ &= \frac{1}{n-k+1} \prod_{i=n-k+2}^n \frac{i-1}{i} \\ &= \frac{1}{n}. \end{aligned}$$

$$E(X) = \sum_{i=1}^n i * \frac{1}{n} = \frac{n+1}{2}$$

$$E(X^2) = \sum_{i=1}^n i^2 * \frac{1}{n} = \frac{(n+1)(2n+1)}{6}$$

$$Var(X) = \frac{(n+1)(2n+1)}{6} - \left(\frac{n+1}{2}\right)^2 = \frac{n^2 - 1}{12}$$