

ECSE 305, W09  
Assignment #6, Solutions

$$1. a) E(X) = \sum_{x=1}^5 x p(x) = \frac{1}{15} \sum_{x=1}^5 x^2 = \frac{55}{15} = \frac{11}{3}$$

$$E(X^2) = \sum_{x=1}^5 x^2 p(x) = \frac{1}{15} \sum_{x=1}^5 x^3 = \frac{225}{15} = 15$$

$$\text{Var}(X) = E(X^2) - E(X)^2 = 15 - \frac{121}{9} = \frac{14}{9}$$

b) From the table, we see that  $R_Y = \{5, 8, 9\}$ .

If  $y \notin R_Y$ , then  $p_Y(y) = 0$

$$\begin{aligned} p_Y(5) &= P(Y=5) \\ &= P(X=1) + P(X=5) \\ &= \frac{1}{15} + \frac{5}{15} = \frac{2}{5} \end{aligned}$$

$$\begin{aligned} p_Y(8) &= P(X=2) + P(X=4) \\ &= \frac{2}{15} + \frac{4}{15} = \frac{2}{5} \end{aligned}$$

$$p_Y(9) = P(X=3) = \frac{1}{5}$$

X	Y
1	5
2	8
3	9
4	8
5	5

$$c) E(Y) = \sum_{y \in R_Y} y p_Y(y) = 5 \cdot \frac{2}{5} + 8 \cdot \frac{2}{5} + 9 \cdot \frac{1}{5} = 7$$

$$\begin{aligned} E(Y) &= E(X(6-X)) = E(6X - X^2) \\ &= 6E(X) - E(X^2) = 6 \cdot \frac{11}{3} - 15 = 7 \end{aligned}$$

$$E(Y) = \sum_{x=1}^5 x(6-x) p(x) = \dots = 7$$

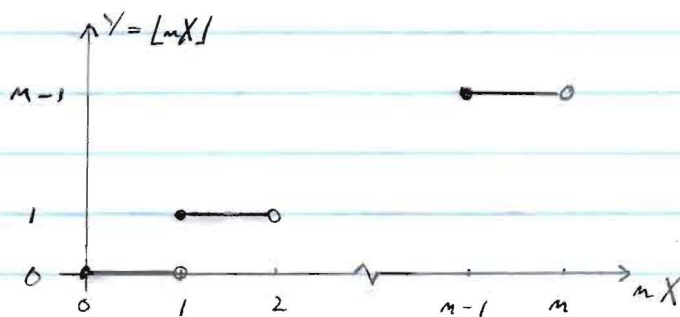
2. Let  $X$  denote the number of errors on the specific page. Then  $X$  is a Poisson random variable with  $E(X) = \lambda = 1/5$ . We want

$$\begin{aligned} P(X \geq 1) &= 1 - P(X=0) \\ &= 1 - e^{-\lambda} \\ &\approx 0.18 \end{aligned}$$

3.  $X$  is a continuous RV with uniform PDF

$$f(x) = \begin{cases} 1, & 0 < x < 1 \\ 0, & \text{otherwise} \end{cases}$$

For  $n$  a positive integer, define  $Y = \lfloor nX \rfloor$ . The situation is illustrated below:



We see that  $Y$  is a discrete RV from the set  $R_Y = \{0, 1, \dots, n-1\}$ . For any  $y \in R_Y$ , we have

$$\begin{aligned} P_Y(y) &= P(Y=y) \\ &= P(y \leq nX < y+1) \\ &= P\left(\frac{y}{n} \leq X < \frac{y+1}{n}\right) = \frac{1}{n} \end{aligned}$$

4. The length of the other side is  $Y = \sqrt{81 - X^2}$ .  
Therefore, the expected value is

$$\begin{aligned}
 E[Y] &= \int_{-\infty}^{\infty} \sqrt{81 - x^2} f(x) dx \\
 &= \frac{1}{6} \int_2^4 \sqrt{81 - x^2} x dx \\
 &= \frac{1}{12} \int_{65}^{77} \sqrt{u} du, \quad \begin{array}{l} u = 81 - x^2 \\ du = -2x dx \end{array} \\
 &= \frac{1}{18} u^{3/2} \Big|_{65}^{77} \\
 &\approx 8.4
 \end{aligned}$$

5. More generally, consider the Laplacian RV (Section 7.5.3) with PDF

$$f(x) = \frac{\alpha}{2} e^{-\alpha|x|}, \quad x \in \mathbb{R}$$

where  $\alpha > 0$  is an arbitrary parameter. Note that

$$\int_{-\infty}^{\infty} f(x) dx = \alpha \int_0^{\infty} e^{-\alpha x} dx = 1$$

Also, by symmetry,  $E(X) = \mu = 0$ . We have

$$\begin{aligned}
 E(X^2) &= \int_{-\infty}^{\infty} x^2 f(x) dx \\
 &= \alpha \int_0^{\infty} x^2 e^{-\alpha x} dx \\
 &\stackrel{(1)}{=} \alpha \left. \frac{e^{-\alpha x}}{(-\alpha)} \left( x^2 + \frac{2x}{\alpha} + \frac{2}{\alpha^2} \right) \right|_0^{\infty}
 \end{aligned}$$

<sup>(1)</sup> This is from Table of integrals; can also do integration by parts -3-

$$= \frac{2}{\alpha^2}$$

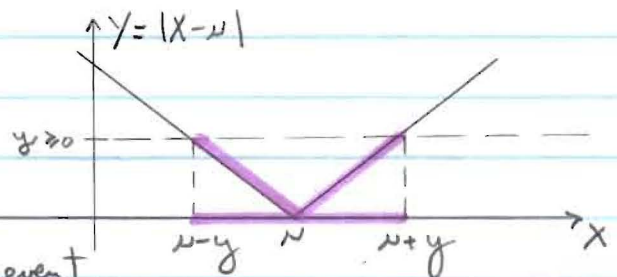
$$\text{Var}(X) = E(X^2) - \mu^2 = \frac{2}{\alpha^2}$$

In the special case of interest here, i.e.  $\alpha=1$ , we find  $\text{Var}(X) = 2$ .

$$\begin{aligned} 6. \quad E[e^X] &= \int_{-\infty}^{\infty} e^x f(x) dx \\ &= 3 \int_0^{\infty} e^{-2x} dx \\ &= 3 \left( \frac{e^{-2x}}{-2} \right) \Big|_0^{\infty} = \frac{3}{2} \end{aligned}$$

7. Different approaches are possible for finding the CDF of  $Y = |X - \mu|$ .

A) Method of distribution:



Step ①:

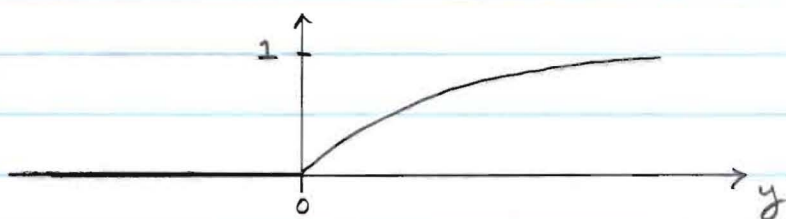
case  $y < 0$ :  $Y \leq y \iff$  impossible event  
 case  $y \geq 0$ :  $Y \leq y \iff \mu - y \leq X \leq \mu + y$

Step ②:

$$\begin{aligned} \text{case } y < 0: \quad G(y) &= P(Y \leq y) = 0 \\ \text{case } y \geq 0: \quad G(y) &= P(\mu - y \leq X \leq \mu + y) \\ &= P(X \leq \mu + y) - P(X < \mu - y) \\ &= \Phi\left(\frac{y}{\sigma}\right) - \Phi\left(-\frac{y}{\sigma}\right) \quad (\text{Note: } \Phi(-x) = 1 - \Phi(x)) \\ &= 2\Phi\left(\frac{y}{\sigma}\right) - 1 \end{aligned}$$

Therefore, the CDF of  $Y$  is

$$G(y) = \begin{cases} 0, & y \leq 0 \\ 2\Phi\left(\frac{y}{\sigma}\right) - 1, & y > 0 \end{cases}$$



B) Method of transformation: For  $y < 0$ ,  $y = |x - \mu|$  has no root and so  $g(y) = 0$ . For  $y > 0$ , there are 2 roots:

$$x_1 = \mu - y, \quad \frac{dx_1}{dy} = -1$$

$$x_2 = \mu + y, \quad \frac{dx_2}{dy} = +1$$

$$g(y) = \sum_{i=1}^2 f(x_i) \left| \frac{dx_i}{dy} \right|$$

$$= f(\mu - y) + f(\mu + y), \quad f(x) = \frac{1}{\sigma} \phi\left(\frac{x - \mu}{\sigma}\right)$$

$$= \frac{1}{\sigma} \phi\left(-\frac{y}{\sigma}\right) + \frac{1}{\sigma} \phi\left(\frac{y}{\sigma}\right), \quad y > 0$$

Finally,  $G(y)$  is obtained by integration:

$$G(y) = \int_{-\infty}^y g(t) dt$$

$$= \begin{cases} 0 & y < 0 \\ \Phi\left(\frac{y}{\sigma}\right) - \Phi\left(-\frac{y}{\sigma}\right) & y \geq 0 \end{cases}$$

For the expected value, we have

$$E[Y] = E[|X - \mu|] = \int_{-\infty}^{\infty} |x - \mu| \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

We make the change of variable  $t \triangleq (x - \mu)/\sigma$ ,  $dt = dx/\sigma$ :

$$\begin{aligned} E[Y] &= \frac{2\sigma}{\sqrt{2\pi}} \int_0^{\infty} t e^{-t^2/2} dt \\ &= \frac{\sqrt{2}}{\sqrt{\pi}} \sigma e^{-t^2/2} \Big|_0^{\infty} \\ &= \frac{\sqrt{2}}{\sqrt{\pi}} \sigma \end{aligned}$$

8. Let  $X$  denote the chest size of a randomly selected soldier. We assume that  $X \sim \mathcal{N}(\mu, \sigma^2)$  with  $\mu = 40$  in and  $\sigma = 2.5$  (the inflection pts are located at  $\mu \pm \sigma$  for a normal distribution). Then

$$\begin{aligned} p &\leq P(X \geq 40) \\ &= P\left(\frac{X - 40}{2.5} \geq \frac{40 - 40}{2.5}\right) \\ &= P(Z \geq 0) = \frac{1}{2} \quad \text{where } Z \triangleq \frac{X - \mu}{\sigma} \sim \mathcal{N}(0, 1) \end{aligned}$$

Now, let  $Y$  denote the number of soldiers, out of 50, with chest size of at least 40. We model  $Y \sim B(n, p)$  with  $n = 50$  and  $p = 1/2$ . We seek

$$P(Y = 10) = \binom{50}{10} (0.5)^{10} (0.5)^{40} \approx 9.12 \times 10^{-6}$$

9. a) Let  $X$  denote the score of a randomly selected student. Then  $X \sim \mathcal{N}(\mu, \sigma^2)$  with  $\mu = 75$  and  $\sigma = 10$ . We seek  $x$  such that  $P(X \geq x) = .1$ . Equivalently:

$$P(X \leq x) = .9 \iff P\left(\frac{X - \mu}{\sigma} < \frac{x - \mu}{\sigma}\right) = .9$$

$$\iff \Phi\left(\frac{x - \mu}{\sigma}\right) = .9$$

since  $Z \triangleq \frac{X - \mu}{\sigma}$  is  $\mathcal{N}(0, 1)$ . From the table for  $\Phi(\cdot)$ , we have that  $\Phi(1.28) \approx .8997$ . Thus

$$\frac{x - \mu}{\sigma} = 1.28$$

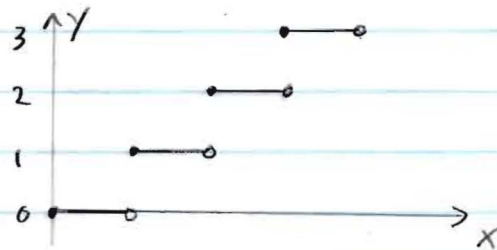
$$\begin{aligned} x &= \mu + 1.28\sigma \\ &= 75 + 12.8 \\ &= 87.8\% \end{aligned}$$

b) We just have to use different values from the table:

$$5\% : \Phi(1.59) \approx .9441 \implies x = 90.9\%$$

$$2\% : \Phi(2.05) \approx .9803 \implies x = 95.6\%$$

10. We note that  $Y$  can only take integer values. Also, for negative integers  $y$ ,  $P(Y = y) = 0$ . For non-negative integers, i.e.  $y \in \mathbb{Q}_y = \{0, 1, 2, \dots\}$ , we have

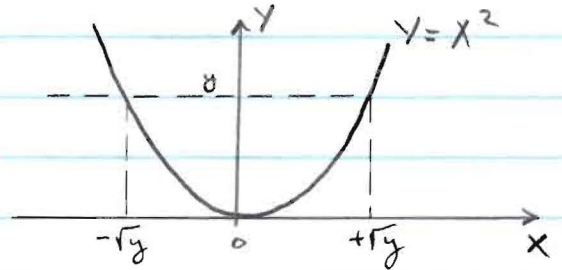


$$p(y) \triangleq P(Y = y) = P(y \leq X < y+1)$$

$$= \int_y^{y+1} \lambda e^{-\lambda x} dx = e^{-\lambda y} (1 - e^{-\lambda})$$

Introducing  $p = 1 - e^{-\lambda}$ , we have  $p q^y = e^{-\lambda y} p$  which corresponds to the PMF of a geometric RV.

11. We can use the method of distribution with  $Y = X^2$ .  
For  $y < 0$ , we have



$$G(y) = P(Y \leq y) \\ = P(X^2 \leq y) = 0$$

$$g(y) = G'(y) = 0$$

For  $y \geq 0$ , we have  $Y \leq y \iff -\sqrt{y} \leq X \leq \sqrt{y}$ . So

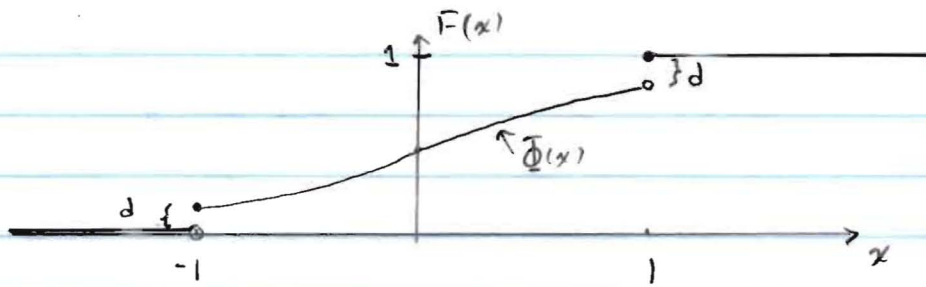
$$G(y) = P(Y \leq y) \\ = P(-\sqrt{y} \leq X \leq \sqrt{y}) \\ = P(X \leq \sqrt{y}) - P(X < -\sqrt{y}) \\ = \Phi(\sqrt{y}) - \Phi(-\sqrt{y}) \\ = 2\Phi(\sqrt{y}) - 1$$

$$g(y) = \frac{1}{\sqrt{y}} \phi(\sqrt{y}) = \frac{1}{\sqrt{y} \sqrt{2\pi}} e^{-y/2} \\ = \frac{1}{\sqrt{\pi}} \frac{1}{2} e^{-y/2} (y/2)^{-1/2} \\ = \frac{\lambda}{\Gamma(\beta)} e^{-\lambda y} (\lambda y)^{\beta-1}$$

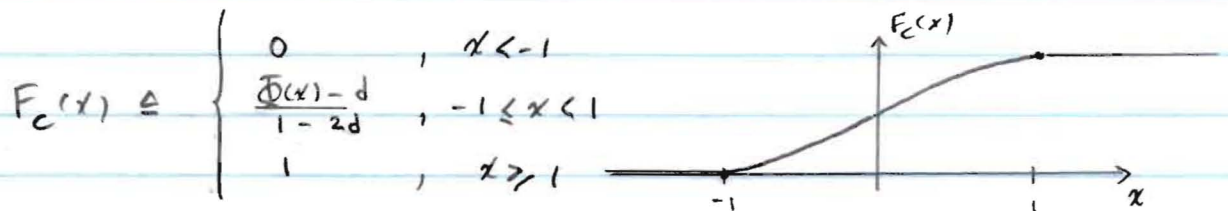
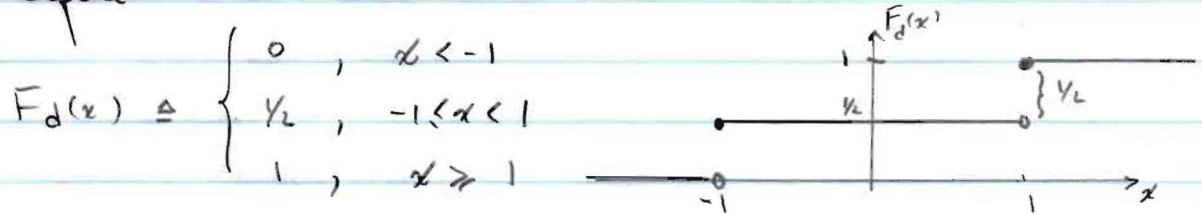
with  $\lambda = \frac{1}{2}$ ,  $\beta = \frac{1}{2}$  and  $\Gamma(\frac{1}{2}) = \sqrt{\pi}$  (see Table of gamma function)



12.



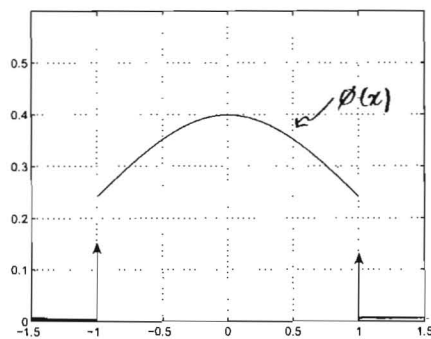
a) It is convenient to define  $d \triangleq \Phi(-1) = 1 - \Phi(1) \approx 0.1587$ .  
 Clearly, we can write  $F(x) = \alpha F_d(x) + \beta F_c(x)$  where we define



Parameter  $\alpha$  is obtained by adjusting the size of the jump at  $x = -1$ :  $\frac{\alpha}{2} = d$ , i.e.  $\alpha = 2d = .3174$  and  $\beta = 1 - \alpha = .6826$

b) From the plot of  $F(x)$  at the top:

$$S(x) = \frac{d}{dx} F(x) = d\delta(x+1) + d\delta(x-1) + \phi(x) [u(x+1) - u(x-1)]$$



$$c) P(X < 0) = F(0^-) = \Phi(0^-) = \Phi(0) = 1/2$$

$$P(X \leq 0) = F(0) = \Phi(0) = 1/2$$

$$P(X < 1) = F(1^-) = \Phi(1^-) = \Phi(1) = .8413$$

$$P(X \leq 1) = F(1) = 1$$

$$d) E(X) = \int_{-\infty}^{\infty} x f(x) dx$$

$$= (-1)d + (1)d + \int_{-1}^1 x \phi(x) dx$$

$$= 0$$

$$\text{Var}(X) = \int_{-\infty}^{\infty} x^2 f(x) dx$$

$$= 2d + \int_{-1}^1 x^2 \phi(x) dx$$

$$= \int_{-\infty}^{-1} \phi(x) dx + \int_1^{\infty} \phi(x) dx + \int_{-1}^1 x^2 \phi(x) dx$$

$$< \int_{-\infty}^{\infty} x^2 \phi(x) dx = 1$$