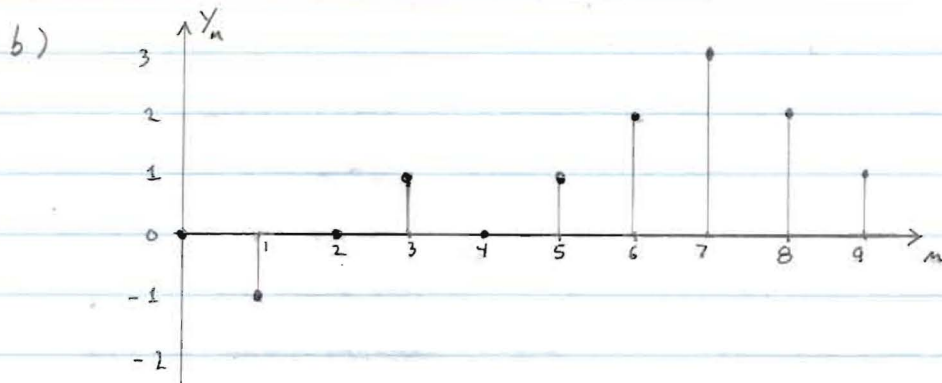


ECSE 305, W09
Assignment #10, Solutions

1. a) • Y_n is discrete-time: time argument n takes value in countably infinite set $T = \{0, 1, 2, \dots\}$.
 • Y_n is discrete-space: the possible values of Y_n are given by the countably infinite set $\Omega = \{0, \pm 1, \pm 2, \dots\}$.



c) We first note that for any $i \geq 1$:

$$\begin{aligned}
 E(X_i) &= (1)p + (-1)q = p - q \\
 E(X_i^2) &= (1)^2 p + (-1)^2 q = p + q = 1 \\
 \text{Var}(X_i) &= E(X_i^2) - E(X_i)^2 \\
 &= 1 - (p - q)^2 = \dots = 4pq
 \end{aligned}$$

Thus, for $n \geq 1$, we have

$$\begin{aligned}
 \mu_Y(n) &\triangleq E(Y_n) = E\left(\sum_{i=1}^n X_i\right) \\
 &= \sum_{i=1}^n E(X_i) = n(p - q)
 \end{aligned}$$

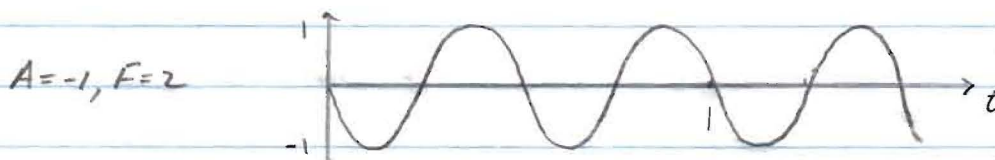
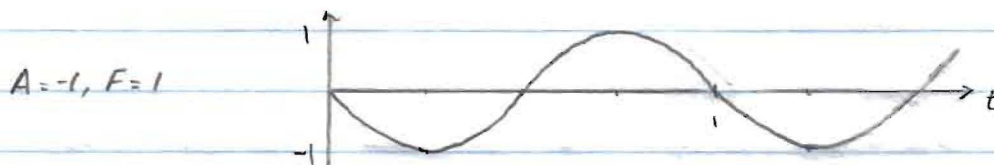
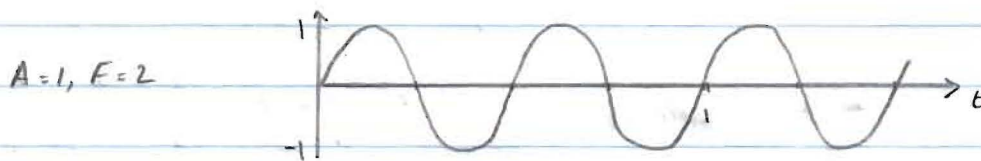
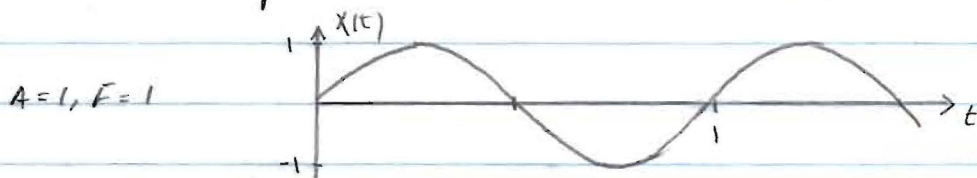
$$\begin{aligned}
 \sigma_Y^2(n) &\triangleq \text{Var}(Y_n) = \text{Var}\left(\sum_{i=1}^n X_i\right) \\
 &= \sum_{i=1}^n \text{Var}(X_i) = 4npq
 \end{aligned}$$

note: the X_i are independent

For $m=0$, $Y_0=0$ and so $E(Y_0)=0$, $\text{Var}(Y_0)=0$. The above formulas also apply to $m=0$.

2. a) • Continuous-time; $t \in T = [0, \infty)$
 • Continuous-space; for an arbitrary t , $X(t) \in \Omega = [-1, 1]$

b) There are 4 possibilities:



c) We first note that $E(A) = (1) \cdot \frac{1}{2} + (-1) \cdot \frac{1}{2} = 0$. Since A and F are independent, we have

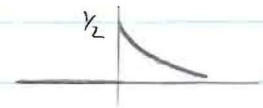
$$\begin{aligned} \mu_X(t) &\triangleq E(X(t)) \\ &= E(A \sin(2\pi Ft)) \\ &= E(A) E(\sin 2\pi Ft) \\ &= 0, \quad \text{all } t \in T \end{aligned}$$

3. a) From Theorem 13.2, we have

$$\begin{aligned}
 R_Y(z) &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} h(u_1) h(u_2) R_X(z - u_1 + u_2) du_1 du_2 \\
 &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} h(u_1) h(u_2) \delta(u_2 - (u_1 - z)) du_1 du_2 \\
 &= \int_{-\infty}^{\infty} h(u_1) h(u_1 - z) du_1 \quad (= h(z) * h(-z))
 \end{aligned}$$

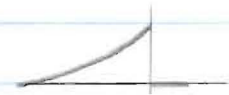
$$\begin{aligned}
 z \geq 0: R_Y(z) &= \int_z^{\infty} e^{-u_1} e^{-(u_1 - z)} du_1 \\
 &= e^z \int_z^{\infty} e^{-2u_1} du_1
 \end{aligned}$$

$$= e^z \left(\frac{e^{-2u_1}}{-2} \right) \Big|_z^{\infty} = \frac{1}{2} e^{-z}$$



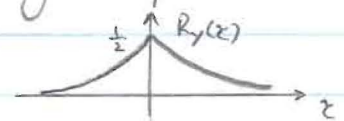
$$z \leq 0: R_Y(z) = \int_0^{\infty} e^{-u_1} e^{-(u_1 - z)} du_1$$

$$= \dots = \frac{1}{2} e^z$$



These two cases can be combined into a single expression:

$$R_Y(z) = \frac{1}{2} e^{-|z|}, \quad z \in \mathbb{R}$$



b) The frequency response of the filter is

$$H(\omega) = \int_{-\infty}^{\infty} h(t) e^{-j\omega t} dt$$

$$= \int_0^{\infty} e^{-(1+j\omega)t} dt = \frac{1}{1+j\omega}$$

The PSD of the input process is

$$S_x(\omega) = \int_{-\infty}^{\infty} R_x(\tau) e^{-j\omega\tau} d\tau$$

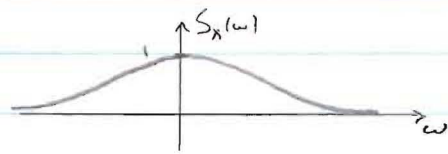
$$= \int_{-\infty}^{\infty} \delta(\tau) e^{-j\omega\tau} d\tau = 1$$

The PSD of the output process is (Theorem 13.3)

$$S_y(\omega) = |H(\omega)|^2 S_x(\omega)$$

$$= \left| \frac{1}{1+j\omega} \right|^2$$

$$= \frac{1}{1+\omega^2}$$



4. $\mu_x(t) \triangleq E[X(t)]$

$$= E[A \cos(\omega t + B)]$$

$$= E[A] E[\cos(\omega t + B)] \quad (\text{because } A \text{ and } B \text{ are independent})$$

$$= 0 \quad (\text{because } A \sim N(0,1))$$

$$R_x(t, u) \triangleq E[X(t)X(u)]$$

$$= E[A^2 \cos(\omega t + B) \cos(\omega u + B)]$$

$$= E[A^2] E[\cos(\omega t + B) \cos(\omega u + B)] \quad (E[A^2] = 1)$$

$$= \frac{1}{2} E[\cos(\omega(t-u)) + \cos(\omega(t+u) + 2B)]$$

$$= \frac{1}{2} \cos(\omega(t-u)) + \frac{1}{2} E[\cos(\omega(t+u) + 2B)]$$

PDF of B

$$= \frac{1}{2} \cos(\omega(t-u)) + \frac{1}{2} \int_0^{2\pi} \cos(\omega(t+u) + 2b) \frac{1}{2\pi} db$$

$$= \frac{1}{2} \cos(\omega(t-u))$$

Since $\mu_x(t)$ is a constant and $R_x(t, u)$ is only a function of the difference (or lag) $t - u$, we conclude that $X(t)$ is WSS.

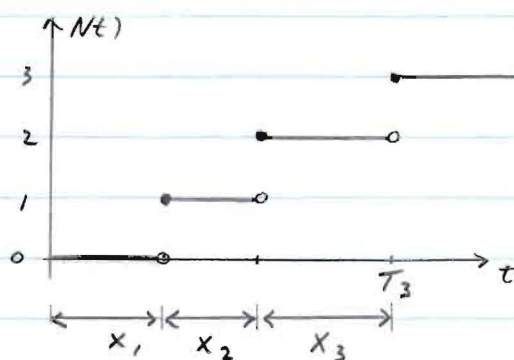
$$5. R_y(t_1, t_2) \triangleq E[Y(t_1)Y(t_2)]$$

$$= E\left[\int_{-\infty}^{\infty} h(u_1) X(t_1 - u_1) du_1, \int_{-\infty}^{\infty} h(u_2) X(t_2 - u_2) du_2\right]$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} h(u_1) h(u_2) E[X(t_1 - u_1) X(t_2 - u_2)] du_1 du_2$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} h(u_1) h(u_2) R_x(\underbrace{(t_1 - u_1) - (t_2 - u_2)}_{= t_1 - t_2 - u_1 + u_2}) du_1 du_2$$

6. a) Let X_i denote the interarrival time of the i th patient. RVs X_i are exponentially distributed with parameter $\lambda = .1 \text{ min}^{-1}$. We seek



$$\begin{aligned} E(T_3) &= E(X_1 + X_2 + X_3) \\ &= E(X_1) + E(X_2) + E(X_3) \\ &= 3 \cdot \frac{1}{\lambda} \\ &= 30 \text{ min} \end{aligned}$$

b) Observe that $T_3 > 60 \text{ min}$ iff $N(60) \leq 2$. Also $N(60)$ is a Poisson RV with parameter $60\lambda = 6$. Hence

$$\begin{aligned} P(T_3 > 60) &= P(N(60) \leq 2) \\ &= \sum_{i=0}^2 P(N(60) = i) \\ &= \sum_{i=0}^2 \frac{6^i}{i!} e^{-6} \\ &= (1 + 6 + 18) e^{-6} \approx 0.062 \end{aligned}$$

7. a) It can be seen that

$$P(X=1, Y=-1) = 0 \neq P(X=1)P(Y=-1) = \frac{1}{4} \cdot \frac{17}{48}$$

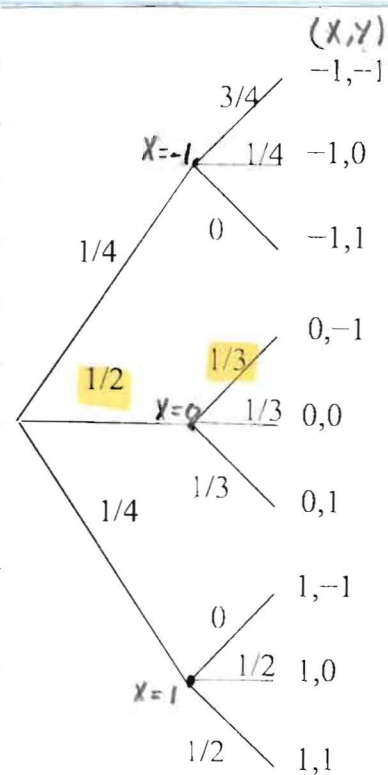
Therefore, X and Y are NOT independent.

b) From the given data, we must first compute the marginal PMF of X , and the conditional PMF of Y given $X=x$, e.g.

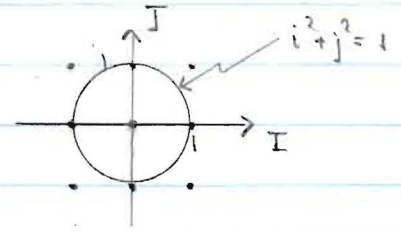
$$P_X(0) = \frac{1}{6} + \frac{1}{6} + \frac{1}{6} = \frac{1}{2}$$

$$P_{Y|X}(-1|0) = \frac{P(0,-1)}{P_X(0)} = \frac{\frac{1}{6}}{\frac{1}{2}} = \frac{1}{3}$$

The resulting tree diagram is shown on the right:



$$\begin{aligned}
 8. a) \quad \sum_{\text{all } (i,j)} p(i,j) &= \sum_{i=-1}^1 \sum_{j=-1}^1 c(i^2+j^2) \\
 &= c(0+4+4 \times 2) \\
 &= 12c \Rightarrow c = \frac{1}{12}
 \end{aligned}$$

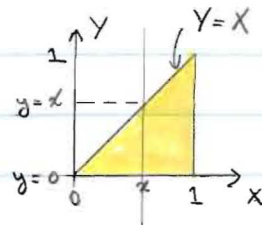


$$\begin{aligned}
 b) \quad p_J(j) &= \sum_{i=-1}^1 p(i,j) \\
 &= \frac{1}{12} [((-1)^2 + j^2) + (0^2 + j^2) + (1^2 + j^2)] \\
 &= \frac{1}{12} (2 + 3j^2), \quad j \in \{-1, 0, 1\}
 \end{aligned}$$

$$\begin{aligned}
 p_{I|J}(i,j) &= \frac{p(i,j)}{p_J(j)} \\
 &= \frac{\frac{1}{12} (i^2 + j^2)}{\frac{1}{12} (2 + 3j^2)} \\
 &= \frac{i^2 + j^2}{2 + 3j^2}, \quad i, j \in \{-1, 0, 1\}
 \end{aligned}$$

$$\begin{aligned}
 c) \quad P(|I|=1 | J=0) &= P(I=-1 | J=0) + P(I=1 | J=0) \\
 &= p_{I|J}(-1/0) + p_{I|J}(1/0) \\
 &= \frac{1}{2} + \frac{1}{2} \\
 &= 1
 \end{aligned}$$

$$9. \quad f(x, y) = \begin{cases} 6y, & 0 \leq y \leq x \leq 1 \\ 0, & \text{otherwise} \end{cases}$$



$$a) \quad f_x(x) = \int_{-\infty}^{\infty} f(x, y) dy$$

$$x \in [0, 1] \Rightarrow f_x(x) = \int_0^x 6y dy = 3x^2$$

$$x \notin [0, 1] \Rightarrow f_x(x) = 0$$

b) $f_{Y|X}(y|x)$ is defined whenever $f_x(x) > 0$, i.e. for $x \in (0, 1]$.
In this case we have

$$f_{Y|X}(y|x) = \frac{f(x, y)}{f_x(x)} = \begin{cases} \frac{2y}{x^2}, & 0 \leq y \leq x \\ 0, & \text{otherwise} \end{cases}$$

c) From part b), we have $f_{Y|X}(y|1/2) = 8y$ for $0 \leq y \leq 1/2$.

$$\begin{aligned} P(Y \leq 1/4 | X = 1/2) &= \int_{-\infty}^{1/4} f_{Y|X}(y|1/2) dy \\ &= \int_0^{1/4} 8y dy \\ &= (4y^2) \Big|_0^{1/4} = \frac{1}{4} \end{aligned}$$

d) $E[Y|X=x]$ is defined for $x \in (0, 1]$:

$$\begin{aligned} E[Y|X=x] &= \int_{-\infty}^{\infty} y f_{Y|X}(y|x) dy \\ &= \int_0^x y \frac{2y}{x^2} dy \\ &= \frac{2}{x^2} \left(\frac{y^3}{3} \right) \Big|_0^x = \frac{2}{3} x \end{aligned}$$