

Question 1.1

Solution:

$$\begin{aligned}
 P(F_2 \cap F_3) &= P((F_2 \cap F_3 \cap F_1) \cup (F_2 \cap F_3 \cap F_1^c)) \\
 &= 0.004 + 0.023 = 0.027 \\
 P(F_1^c) &= P((F_1^c \cap (F_2 \cap F_3)) \cup (F_1^c \cap (F_2 \cap F_3)^c)) \\
 &= P(F_1^c \cap (F_2 \cap F_3)) + P(F_1^c \cap (F_2 \cap F_3)^c) \\
 &= 0.023 + 0.011 = 0.034 \\
 P(F_1) &= 1 - P(F_1^c) = 1 - 0.034 = 0.966
 \end{aligned}$$

Since all information given about F_2 or F_3 appears in probabilities involving $(F_2 \cap F_3), (F_2 \cap F_3)^c, F_1, F_1^c$, it is not possible to separately compute $P(F_2)$ or $P(F_3)$.

Question 1.2

Solution:

$$S = \{H, TH, T^2H, \dots, T^{N-1}H, \dots\}$$

(a)

i) The events $H, TH, T^2H, \dots, T^{N-1}H, \dots \subset S$ are disjoint since T^jH occurs only if $T^iH, i < j$, does not occur and then $T^kH, k > j$ cannot occur.

ii) No. Since the first union is sample space S and the second set $E(S)$ is the set of all subsets of S .

(b) BE (Branding Experiment) stops at the outcome of n-th toss if and only if $T^{n-1}H$ occurs, where $P(T^{n-1}H) = P(T)^{n-1}P(H) = (\frac{3}{4})^{n-1}\frac{1}{4}$ (by independence) Hence BE stops at the outcome of the n-th toss, $n \leq N$, if $E^N = H \cup TH \cup \dots \cup T^{N-1}H$ occurs, By axiom 3, E^N has probability $P(E^N) = \frac{1}{4} + (\frac{3}{4})\frac{1}{4} + \dots + (\frac{3}{4})^{N-1}\frac{1}{4} = \frac{1}{4} \times \frac{1 - (\frac{3}{4})^N}{1 - \frac{3}{4}} = 1 - (\frac{3}{4})^N$

(c) The events $\{H, TH, T^2H, \dots, T^{N-1}H, \dots\}$ are not pairwise independent since they are disjoint,

or events $T^iH, T^jH, 0 \leq i < j < \infty$ are not independent since $0 = P(T^iH \cap T^jH) \neq P(T^iH)P(T^jH) = (\frac{3}{4})^{i-1}\frac{1}{4}(\frac{3}{4})^{j-1}\frac{1}{4}, 0 \leq i < j < \infty$.

(d)

i) The probability that the experiment BE stops after an odd number of tosses is:

$$\sum_{k=0}^{\infty} \frac{1}{4} \left(\frac{3}{4}\right)^{2k} = \frac{1}{4} \times \frac{1}{1 - \left(\frac{3}{4}\right)^2} = \frac{4}{7}.$$

ii) Let $A = \{\text{The experiment BE stops after an even number of tosses}\}$

and $B = \{\text{The experiment BE stops after an odd number of tosses}\}$ then we have

$$P(S) = P(A) + P(B) = 1.$$

Therefore, $P(A) = 1 - P(B) = 1 - \frac{4}{7} = \frac{3}{7}$.

Question 1.3

Solution:

(a) The sample space for this experiment where the outcome is taken to be pair (T_1, T_2) is shown in Fig. 1.

(b) The region A of the plane corresponding to the event “student is awake at 9pm” is shown in Fig. 2.

(c) The set B in the plane corresponding to the event “student is asleep more time than the student is awake” is shown in Fig. 3.

(d) The region corresponding to the event $D = A^c \cap B$, which means “student is not awake at 9pm and the student is asleep more time than being awake” is shown in Fig. 4.

The probability of the event D is

$$P_D = \frac{\frac{1}{2} \times 15^2 - \frac{1}{2} \times 3^2 + \frac{1}{2} \times 9^2}{\frac{1}{2} \times 24^2} \approx 0.516$$

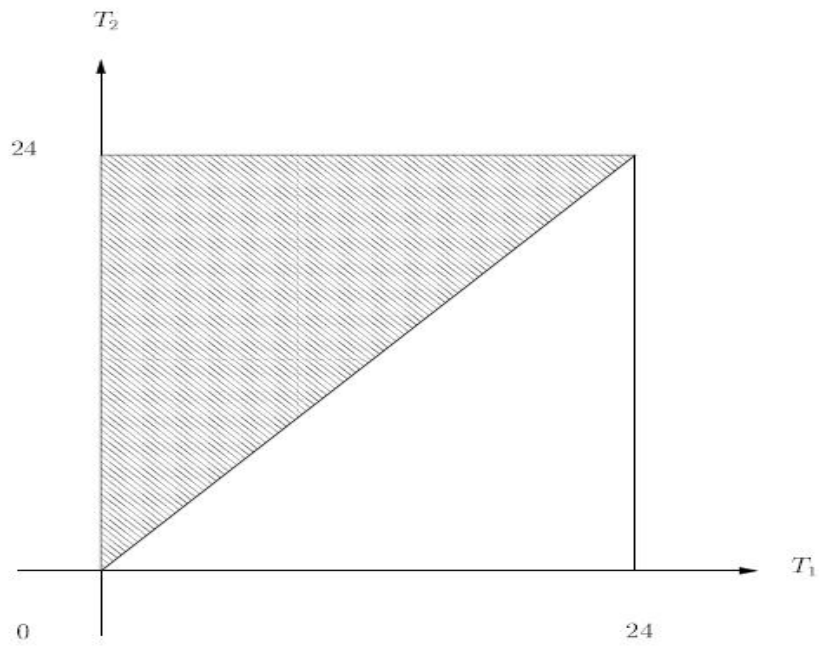


Figure 1: Question 1.3 (a): sample space S .

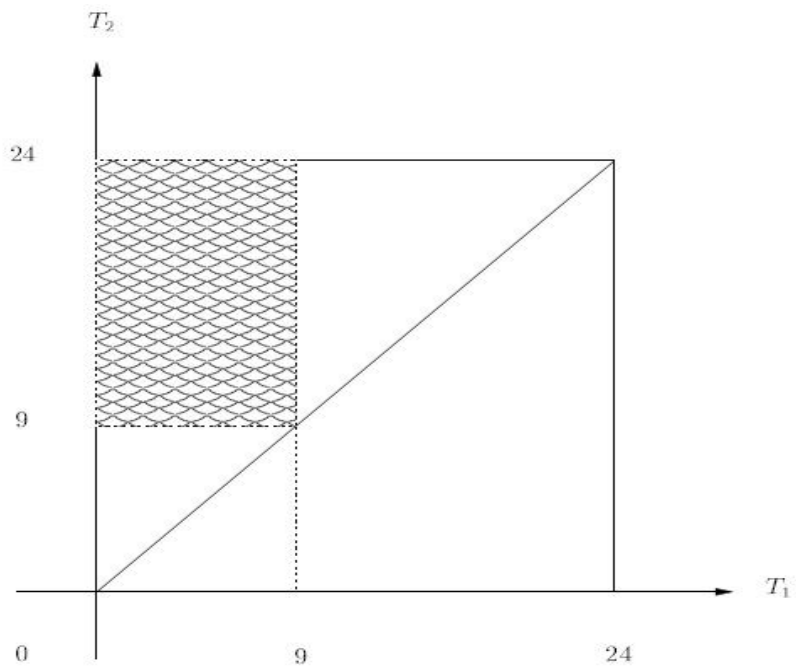


Figure 2: Question 1.3 (b): region corresponding to the event A .

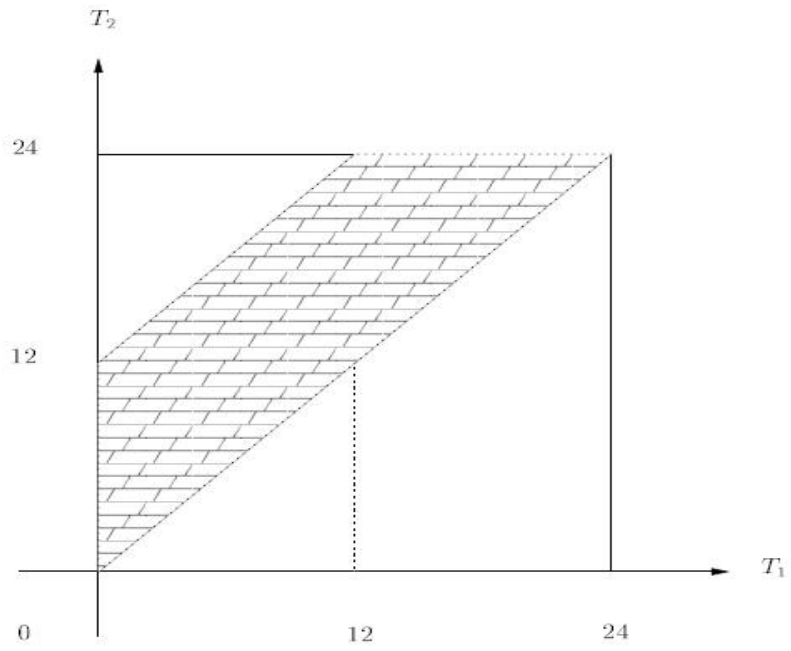


Figure 3: Question 1.3 (c): region corresponding to the event B .

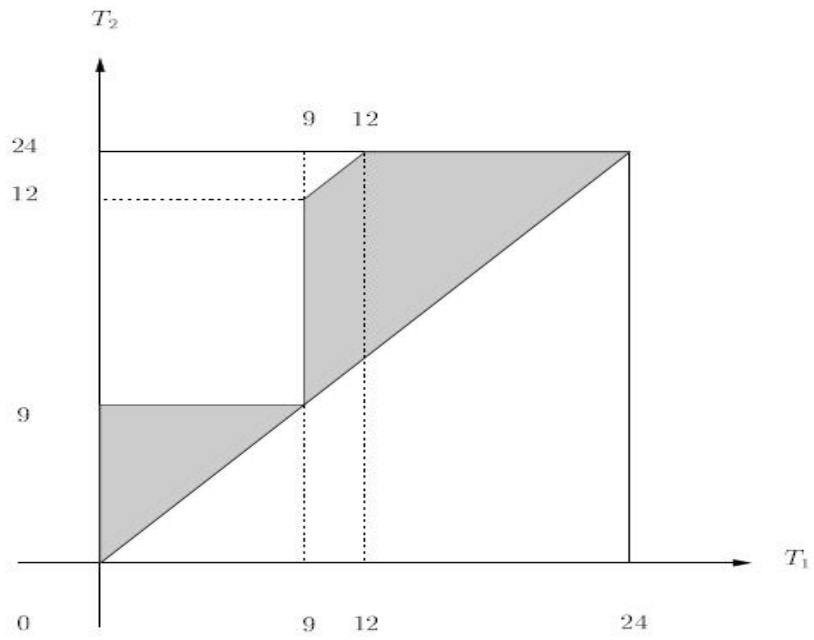


Figure 4: Question 1.3 (d): region corresponding to the event $D = A^c \cap B$.

Question 1.4

Solution:

$$\text{i) } P(A^c) = \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \text{Ind}_k(A^c) = \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n [1 - \text{Ind}_k(A)] = 1 - [\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \text{Ind}_k(A)] = 1 - P(A).$$

ii) $\text{Ind}_k(A) = 0$ or 1 , and $\text{Ind}_k(B) = 0$ or 1 , so

$$0 \leq \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \text{Ind}_k(A) \leq 1,$$

$$0 \leq \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \text{Ind}_k(B) \leq 1.$$

Since

$$\text{Ind}_k(A) = 1 \Rightarrow \text{Ind}_k(B) = 1,$$

$$\text{Ind}_k(B) = 1 \Rightarrow \text{Ind}_k(A) = 0 \text{ or } 1,$$

then

$$\frac{1}{n} \sum_{k=1}^n \text{Ind}_k(A) \leq \frac{1}{n} \sum_{k=1}^n \text{Ind}_k(B),$$

and so

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \text{Ind}_k(A) \leq \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \text{Ind}_k(B),$$

therefore $0 \leq P(A) \leq P(B) \leq 1$.

iii) Let $(AB)_k$, etc., denote $(A \cap B)_k$, etc., then

$$\text{Ind}_k(A \cup B) \begin{Bmatrix} 1 & \text{on } (AB)_k, \\ 1 & (AB^c)_k, \\ 1 & (A^c B)_k, \\ 0 & (A^c B^c)_k. \end{Bmatrix} = \text{Ind}_k(A) \begin{Bmatrix} 1 & \text{on } (AB)_k, \\ 1 & (AB^c)_k, \\ 0 & (A^c B)_k, \\ 0 & (A^c B^c)_k. \end{Bmatrix} + \text{Ind}_k(B) \begin{Bmatrix} 1 & \text{on } (AB)_k, \\ 0 & (AB^c)_k, \\ 1 & (A^c B)_k, \\ 0 & (A^c B^c)_k. \end{Bmatrix} - \text{Ind}_k(A \cap B) \begin{Bmatrix} 1 & \text{on } (AB)_k, \\ 0 & (AB^c)_k, \\ 0 & (A^c B)_k, \\ 0 & (A^c B^c)_k. \end{Bmatrix}.$$

So

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \text{Ind}_k(A \cup B) = \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \text{Ind}_k(A) + \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \text{Ind}_k(B) - \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \text{Ind}_k(A \cap B),$$

therefore

$$P(A \cup B) = P(A) + P(B) - P(A \cap B).$$

We note that in the calculation above, $\text{Ind}_k(A \cap B) = \text{Ind}_k(A) \cdot \text{Ind}_k(B)$.

Question 1.5

Solution:

Define the events W_I , W_{II} , and W_{III} as the candidate winning districts I, II and III separately; we seek the probability $P(W_I \cap W_{II} \cap W_{III})$. It is known that $P(W_I \cap W_{III}) = 0.55$, $P(W_{II}^c \cap W_I) = 0.35$, and $P(W_{II}^c \cap W_{III} \cap W_I) = 0.15$.

Hence

$$\begin{aligned} P(W_I \cap W_{II} \cap W_{III}) &= P(W_I \cap W_{III}) - P(W_I \cap W_{II}^c \cap W_{III}) \\ &= P(W_I \cap W_{III}) - [P(W_I \cap W_{II}^c) - P(W_I \cap W_{II}^c \cap W_{III}^c)] \\ &= 0.55 - 0.35 + 0.15 \\ &= 0.35. \end{aligned}$$