

**7.1 Solution:**

(i) Since  $X_i$  are independent random variables and uniformly distributed on the interval  $[-k, k]$ , the characteristic functions are given by

$$\Phi_{X_i}(w) = \int_{-k}^k e^{jwx} \frac{1}{2k} dx = \frac{e^{jwk} - e^{-jwk}}{2kjw},$$

which are even functions since  $\Phi_{X_i}(w) = \Phi_{X_i}(-w)$ .

Given  $Y_n = \sum_{k=1}^n X_k$ ,

$$\begin{aligned} \Phi_{Y_n}(w) &= E(\exp^{jwY_n}) = E(\exp^{jw(X_1+X_2+\dots+X_n)}) \\ &= E(\exp^{jwX_1})E(\exp^{jwX_2}) \dots E(\exp^{jwX_n}) \\ &= \Phi_{X_1}(w)\Phi_{X_2}(w) \dots \Phi_{X_n}(w). \end{aligned}$$

Given  $Z_n = \sum_{k=1}^n (-1)^k X_k$ ,

$$\begin{aligned} \Phi_{Z_n}(w) &= E(\exp^{jwZ_n}) = E(\exp^{jw(-X_1+X_2+\dots+(-1)^n X_n)}) \\ &= E(\exp^{-jwX_1})E(\exp^{jwX_2}) \dots E(\exp^{(-1)^n jwX_n}) \\ &= \Phi_{X_1}(-w)\Phi_{X_2}(w) \dots \Phi_{X_n}((-1)^n w) \end{aligned}$$

Since each  $\Phi_i$  is an even function of  $w$ , we have

$$\Phi_{Z_n}(w) = \Phi_{X_1}(w)\Phi_{X_2}(w) \dots \Phi_{X_n}(w),$$

that is to say  $\Phi_{Z_n}(w) = \Phi_{Y_n}(w)$ .

And because distribution functions are in one to one relation with characteristic functions, we conclude that  $Y_n$  and  $Z_n$  have identical distributions.

(ii) As  $e^x = \lim_{n \rightarrow \infty} (1 + \frac{x}{n})^n$  we have

$$\Phi_{\infty}(w) = e^{\mu jw - \frac{w^2 \sigma^2}{2}}.$$

(iii) It is a Gaussian distribution function whereas

$$f_{\infty}(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}.$$

## 7.2 Solution:

(i)

$$\Phi_{T_\lambda}(w) = \int_{-\infty}^{\infty} e^{jwt} \cdot f_t(t) dt = \int_0^{\infty} e^{jwt} \cdot \lambda \cdot e^{-\lambda t} dt = \frac{\lambda}{\lambda - jw}.$$

(ii) Similar to part (i), we have

$$\Phi_{T_\mu}(w) = \frac{\mu}{\mu - jw}.$$

Since  $T_\lambda$  and  $T_\mu$  are independent random variables, the characteristic function of the total waiting time  $T_\lambda + T_\mu$  is

$$\Phi_{T_\lambda + T_\mu}(w) = \Phi_{T_\lambda}(w) \cdot \Phi_{T_\mu}(w) = \frac{\lambda\mu}{(\lambda - jw)(\mu - jw)}.$$

(iii)

$$\Phi_{2T_\lambda}(w) := E[e^{jw2T_\lambda}] = E[e^{j(2w)T_\lambda}] = \Phi_{T_\lambda}(2w) = \frac{\lambda}{\lambda - j(2w)}.$$

(iv)

$$\begin{aligned} \Phi_{T_\lambda + T_\mu}(w) \Big|_{\mu=\lambda} &= \frac{\lambda\mu}{(\lambda - jw)(\mu - jw)} \Big|_{\mu=\lambda} = \frac{\lambda^2}{(\lambda - jw)^2} \\ &= \frac{\lambda^2}{\lambda^2 - 2j\lambda\mu - w^2} \neq \Phi_{2T_\lambda}(w). \end{aligned}$$

Hence their density functions could not be the same, since

$$f_x(x) = (1/2\pi) \int_{-\infty}^{\infty} \Phi_x(w) \cdot \exp(-jwx) dw.$$

(v)

$$\begin{aligned} EX^2 &= \frac{1}{j^2} \frac{d^2}{dw^2} \Phi_x(w) \Big|_{w=0} \\ &= \frac{2}{\mu^2} + \frac{2}{\lambda^2} + \frac{2}{\mu\lambda}. \end{aligned}$$

### 7.3 Solution:

(i) By CLT,

$$\frac{1}{\sqrt{n}} \sum_{i=1}^n \frac{X_i}{\sqrt{4}} \approx N(0, 1)$$

Therefore,

$$\begin{aligned} P(-\alpha \leq \frac{1}{n} \sum_{i=1}^n X_i \leq \alpha) &= P(-\frac{\alpha\sqrt{n}}{2} \leq \frac{1}{\sqrt{4n}} \sum_{i=1}^n X_i \leq \frac{\alpha\sqrt{n}}{2}) \\ &\approx \Phi(\frac{\alpha\sqrt{n}}{2}) - \Phi(-\frac{\alpha\sqrt{n}}{2}) \\ &= 2\Phi(\frac{\alpha\sqrt{n}}{2}) - 1 \end{aligned}$$

(ii) When  $\alpha=1$ :

$$(a) 2\Phi(\frac{\sqrt{n}}{2}) - 1 \geq 0.95 \Rightarrow \Phi(\frac{\sqrt{n}}{2}) \geq 0.975 \Rightarrow \frac{\sqrt{n}}{2} \geq 1.96 \Rightarrow n \geq 15.37$$

Since  $n$  is an integer, it has to be equal to or bigger than 16.

$$(b) 2\Phi(\frac{\sqrt{n}}{2}) - 1 \geq 0.9786 \Rightarrow \Phi(\frac{\sqrt{n}}{2}) \geq 0.9893 \Rightarrow \frac{\sqrt{n}}{2} \geq 2.3 \Rightarrow n \geq 21.16$$

Since  $n$  is an integer, it has to be equal to or bigger than 22.

### 7.4 Solution:

(a)

$$\begin{aligned} \Phi_Z(\omega) &= \int_{-\infty}^{\infty} f_Z(z) e^{j\omega z} dz = \int_0^{\infty} 5\mu e^{(j\omega - 5\mu z)} dz \\ &= \frac{5\mu}{j\omega - 5\mu} \cdot e^{z(j\omega - 5\mu)} \Big|_0^{\infty} = \frac{5\mu}{j\omega - 5\mu} \end{aligned}$$

(b)

$$\begin{aligned} \Phi_X(\omega) &= \int_{-\infty}^{\infty} f_X \cdot e^{j\omega x} dx = \int_{-\infty}^0 \frac{\lambda}{2} e^{x(j\omega + \lambda)} dx + \int_0^{\infty} \frac{\lambda}{2} e^{x(j\omega - \lambda)} dx \\ &= \frac{\lambda}{2(j\omega + \delta)} + \frac{\lambda}{2(j\omega - \lambda)} = \frac{\lambda^2}{\omega^2 + \lambda^2} \end{aligned}$$

(c) From the Fourier transform we can get  $\Phi_{-Z}(\omega) = \Phi_Z(-\omega)$

As  $Z_1$  and  $Z_2$  are independent  $\Phi_W(\omega) = \Phi_{Z_1}(\omega) \cdot \Phi_{Z_2}(\omega) = \frac{\lambda^2}{\lambda^2 + \omega^2}$

(d) From the one-to-one relationship, and the characteristic function got above, the probability density of  $W$  is  $f_W(w) = \frac{\lambda}{2} e^{-\lambda|w|}$